

1 Drivers and dynamics of a massive adaptive 2 radiation in cichlid fishes

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12 **Adaptive radiation is the likely source of much of the ecological and morphological diversity of**
13 **life¹⁻⁴. How adaptive radiations proceed and what determines their extent remains elusive in**
14 **most cases^{1,4}. Here we report the in-depth examination of the spectacular adaptive radiation of**
15 **cichlid fishes in African Lake Tanganyika. Based on whole-genome phylogenetic analyses,**
16 **multivariate morphological measurements of three ecologically relevant trait complexes (body**
17 **shape, upper oral jaw morphology, and lower pharyngeal jaw shape), scoring of pigmentation**
18 **patterns, and approximations of the ecology of virtually all ~240 cichlid species endemic to Lake**
19 **Tanganyika, we show that the radiation occurred within the confines of the lake and that**
20 **morphological diversification proceeded in consecutive trait-specific pulses of morphospace**
21 **expansion. We provide empirical support for two theoretical predictions on how adaptive**
22 **radiations proceed, the ‘early-burst’ scenario^{1,5} (for body shape) and the stages model^{1,6,7} (for**
23 **all traits investigated). Through the analysis of two genomes per species and by taking**
24 **advantage of the uneven distribution of species in subclades of the radiation, we further show**
25 **that species richness scales positively with per individual heterozygosity, but is not correlated**
26 **with transposable element content, number of gene duplications, or genome-wide levels of**
27 **selection in coding sequences.**

28 At the macroevolutionary level, the diversity of life has mainly been shaped by two antagonistic
29 processes: evolutionary radiations increase and extinction events decrease organismal diversity over
30 time^{5,8,9}. Evolutionary radiations are referred to as adaptive radiations if new lifeforms evolve rapidly
31 through adaptive diversification into a variety of ecological niches, which typically presupposes
32 ecological opportunity^{1-3,10}. Whether or not an adaptive radiation unfolds depends on a variety of
33 extrinsic and intrinsic factors as well as on contingency, whereas the magnitude of an adaptive
34 radiation is determined by the interplay between its main components, speciation (minus extinction)
35 and adaptation to distinct ecological niches^{1,2,4,11}. Despite considerable scientific interest in the
36 phenomenon of adaptive radiation as cradle of organismal diversity^{1,2,10,12,13}, many predictions

37 regarding its drivers and dynamics remain untested, particularly in exceptionally species-rich
38 instances. Here, we examine what some consider as the ‘most outstanding example of adaptive
39 radiation’¹⁴, the species flock of cichlid fishes in African Lake Tanganyika. This cichlid assemblage
40 comprises about 240 species¹⁵, which together feature an extraordinary degree of morphological,
41 ecological, and behavioural diversity^{14–17}. We construct a species tree of Lake Tanganyika’s cichlid
42 fauna based on genome-wide data, demonstrate the adaptive nature of the radiation, reconstruct eco-
43 morphological diversification along the species tree, and test general and cichlid-specific predictions
44 related to adaptive radiation.

45 ***In situ* radiation in Lake Tanganyika**

46 To establish the phylogenetic context of cichlid evolution in Lake Tanganyika, we estimated the age
47 of the radiation through divergence time analyses based on cichlid and other teleost fossils¹⁸, and
48 constructed time-calibrated species trees using 547 newly sequenced cichlid genomes (Extended Data
49 Table 1). Our new phylogenetic hypotheses (Fig. 1, Extended Data Figs. 1-3) support the assignment
50 of the Tanganyikan cichlid fauna into 16 subclades – corresponding to the taxonomic grouping of
51 species into tribes¹⁵ – and confirm that the Tanganyikan representatives of the tribes Coptodonini,
52 Oreochromini, and Tylochromini belong to more ancestral and widespread lineages that have
53 colonised the lake secondarily^{12,15,19} (Supplementary Discussion). It has been under debate whether
54 all endemic Tanganyikan cichlid tribes evolved within the confines of Lake Tanganyika or whether
55 some of them evolved elsewhere before the formation of the lake^{20–22}. Our new time calibrations
56 establish that the most recent common ancestor of the cichlid radiation in Lake Tanganyika lived
57 around 9.7 (±0.5) Ma (Fig. 1), which coincides with the appearance of lacustrine conditions in the
58 Tanganyikan Rift²³. This suggests that the radiation commenced shortly after the lake had formed
59 and that all endemic cichlid tribes have evolved and diversified *in situ*, that is, within the temporal
60 and geographic context of Lake Tanganyika.

61 **Phenotypes correlate with environments**

62 Because – in the case of adaptive radiation – diversification occurs via niche specialisation, a strong
63 association is expected in the extant fauna between the environment occupied by a species and the
64 specific morphological features used to exploit it^{2,3}. To quantify eco-morphological diversification
65 across the radiation, we investigated three trait complexes through landmark-based morphometric
66 analyses. Specifically, we quantified body shape and upper oral jaw morphology using 2D-landmarks
67 acquired from X-ray images and the shape of the lower pharyngeal jaw bone based on 3D-landmarks
68 derived from micro-computed tomography (μCT) scans (Extended Data Fig. 4). To approximate the
69 ecological niche of each species we used the carbon and nitrogen stable isotope composition of
70 muscle tissue, which informs about the relative position along the benthic-pelagic axis ($\delta^{13}\text{C}$ value)
71 and the relative trophic level ($\delta^{15}\text{N}$ value), respectively^{16,24} – a pattern which we corroborate here for
72 Lake Tanganyika (Extended Data Fig. 5, Supplementary Discussion). The major axes of shape
73 variation for each trait complex were identified through a principal component analysis (PCA). To
74 test for phenotype-environment correlations and to identify the ecologically most relevant
75 components of each of these trait complexes, we performed a two-block partial least square analysis
76 (PLS) with the stable isotope measurements, and applied a phylogenetic generalised least square
77 analysis (pGLS) to account for phylogenetic dependence.

78 The quantification of variation in body shape revealed that PC1 mainly represented
79 differences in aspect ratio, while PC2 was loaded with changes in head morphology (Fig. 2a). The
80 changes in aspect ratio (comparable to PC1) are correlated with the $\delta^{13}\text{C}$ and $\delta^{15}\text{N}$ values (PLS:
81 Pearson's $r = 0.69$, $R^2 = 0.48$, $P = 0.001$; pGLS: $R^2 = 0.12$, $P < 0.001$, $\lambda_{\text{pGLS}} = 1.007$). PC1 of upper
82 oral jaw morphology mainly represented changes in the orientation and relative size of the premaxilla,
83 which was also the main correlate to the stable C and N isotope composition (PLS: Pearson's $r =$
84 0.62 , $R^2 = 0.38$, $P = 0.001$; pGLS: $R^2 = 0.09$, $P < 0.001$, $\lambda_{\text{pGLS}} = 1.023$), while PC2 was defined by
85 changes in the ratio of the rostral versus the lateral part of the bone (Fig. 2b). For lower pharyngeal
86 jaw shape we found that PC1 mainly reflected changes in the aspect ratio of the jaw bone in
87 combination with an increased posterior thickness, while PC2 involved similar shifts in thickness, yet
88 in this case in combination with changes in the lengths of the postero-lateral horns that act as muscle
89 attachment structures²⁵ (Fig. 2c). The PLS revealed that shape changes similar to PC2 are best
90 associated with stable isotope values (PLS: Pearson's $r = 0.67$, $R^2 = 0.45$, $P = 0.001$; pGLS: $R^2 = 0.16$,
91 $P < 0.001$, $\lambda_{\text{pGLS}} = 1.018$). The PCAs further revealed that the occupied area of the morphospace and
92 ecospace scales with the number of species in the subclades (Extended Data Fig. 5; body: Pearson's
93 $r = 0.91$, $df = 9$, $P < 0.001$; oral jaw: Pearson's $r = 0.88$, $df = 9$, $P < 0.001$; pharyngeal jaw: Pearson's
94 $r = 0.83$, $df = 9$, $P = 0.002$; ecospace: Pearson's $r = 0.88$, $df = 9$, $P < 0.001$) – a pattern which is not
95 only driven by sample size (Supplementary Discussion).

96 Overall, the significant association between each of the three traits and the stable C and N
97 isotope composition underpins their adaptive value (Extended Data Fig. 6). A joint consideration
98 points out that deep-bodied cichlids with an inferior mouth and thick lower pharyngeal jaws with
99 short horns are associated with higher stable isotope projections (high $\delta^{13}\text{C}$ and low $\delta^{15}\text{N}$ values),
100 indicating that such fishes predominantly occur in the benthic/littoral zone and feed on plants and
101 algae, while more elongated species with a more superior mouth and longer and thinner lower
102 pharyngeal jaws are generally associated with lower stable isotope projections (low $\delta^{13}\text{C}$ and high
103 $\delta^{15}\text{N}$ values), suggesting a more pelagic lifestyle and a higher position in the food chain.

104 **Pulses of morphological diversification**

105 Next, we investigated the temporal dynamics of how the observed eco-morphological disparity
106 emerged over the course of the radiation. In addition to the three eco-morphological traits, we also
107 scored male pigmentation patterns to approximate disparity along the signalling axis – another
108 potentially important component of diversification in adaptive radiations^{1,6,7,26}. For all four traits, we
109 estimated morphospace expansion through time using ancestral state reconstructions along the time-
110 calibrated species tree and applying a variable rate model of trait evolution^{27,28}. We calculated
111 morphological disparity as extent of occupied morphospace in time intervals of 0.15 million years in
112 comparison to a null model that assumes Brownian motion. Likewise, evolutionary rates through time
113 were calculated as mean evolutionary rates derived from the variable rates model sampled at the same
114 timepoints along the phylogeny.

115 Our analyses uncovered a pattern of discrete pulses in morphospace expansion, which were
116 followed, in most cases, by morphospace packing (Fig. 3). Importantly, the timing of these pulses
117 differed among the traits. For body shape, we found a pulse of rapid morphospace expansion early in
118 the radiation, alongside with the first pulse of lower pharyngeal jaw shape diversification (Fig. 3b,
119 c); this early phase of the radiation also features the highest evolutionary rates for body shape (Fig.
120 3d). The pulse in upper oral jaw diversification occurred in the middle phase of the radiation.

121 Evolutionary rates were elevated during this period, yet even higher at a later phase that was
122 dominated by packing of the upper oral jaw morphospace rather than its expansion (Fig. 3b-d). This
123 suggests that – in that later phase – rapidly evolving lineages diverged into pre-occupied regions of
124 the morphospace, ultimately resulting in convergent forms¹⁶. The second pulse in lower pharyngeal
125 jaw morphospace expansion happened late in the radiation when also evolutionary rates were highest
126 for this trait (Fig. 3b-d). Thus, the theoretical prediction that eco-morphological diversification is
127 rapid early in an adaptive radiation and slows down through time as the available niche space becomes
128 filled^{1,5} applies only to body shape. Yet, this ‘early burst’ in body shape diversification was not
129 connected to a substantial increase in lineage accumulation (Fig. 3c, d).

130 Interestingly, the pigmentation patterns showed a single pulse of diversification and increased
131 evolutionary rates late in the radiation – a signature unlikely to be caused by a high turnover rate in
132 this trait (Supplementary Discussion). This late pulse of diversification in pigmentation patterns,
133 together with the consecutive pulses of morphospace expansion in the eco-morphological traits, is in
134 agreement with the prediction that diversification in an adaptive radiation proceeds in discrete
135 temporal stages – first in macrohabitat use, then by trophic specialisation, followed by a final stage
136 of divergence along the signalling axes^{1,6,7}. However, in contrast to the conventional stages model,
137 the most recent stage of the cichlid adaptive radiation in Lake Tanganyika, which coincides with a
138 large number of speciation events (Fig. 3c, d), is characterised by temporally overlapping pulses of
139 diversification in both a putative signalling trait and in an ecologically relevant trait. The lower
140 pharyngeal jaw shape is the only trait complex showing two discrete pulses of morphospace
141 expansion – one early in the radiation and one late when niche space already became limited. This
142 later pulse suggests that diversification in the pharyngeal jaw apparatus facilitated fine-scaled
143 resource partitioning after body shape and upper oral jaw morphospaces had been explored, resulting
144 in the densely packed niche space observed today (Fig. 3b).

145 **Genomic features and species richness**

146 Finally, we examined whether the diversity patterns arising over the course of the radiation are linked
147 with particular genomic features. It has previously been suggested – based on five reference cichlid
148 genomes – that the radiating African cichlid lineages are characterised by elevated transposable
149 element counts, increased levels of gene duplications, and genome-wide accelerated coding sequence
150 evolution¹³. Because of the phylogenetic sub-structure of Lake Tanganyika’s cichlid fauna and the
151 widely differing species numbers among tribes, our data offered the opportunity to examine genomic
152 features for an association with per-tribe species richness within a large-scale radiation. We did not
153 find evidence that members of species-rich tribes exhibit greater numbers of transposable elements
154 (Fig. 4a) or more duplicated genes in their genomes (Fig. 4b), nor do they feature elevated genome-
155 wide signatures of selection in coding sequences (Fig. 4c). However, we found that a tribe’s species
156 richness scales positively with a common measure of genetic diversity, genome-wide heterozygosity
157 (Fig. 4d). That genetic diversity is linked to species richness has been suspected before, although the
158 nature of this relationship as well as the determinants of genetic diversity are under debate^{29,30}.

159 Elevated levels of heterozygosity could potentially result from hybridisation³¹, which by itself
160 has been suggested as a trigger of cichlid radiations^{22,32,33}. In Tanganyikan cichlids, the level of gene
161 flow within tribes (estimated using f_4 -ratio values³⁴) does not correlate with a tribe’s species richness
162 (Fig. 4e; Extended Data Fig. 8). Nevertheless, much of the variation in heterozygosity as well as its
163 correlation with species richness can be explained by the observed levels of gene flow within tribes

164 in combination with the reduced gene flow among them: Through coalescent simulations of genome
165 evolution along the species tree we show that variation in migration rates, sampled from the empirical
166 f_4 -ratio estimates, can produce levels of heterozygosity that are similar to the ones observed in nature
167 (Fig. 4f). Hence, the correlation between species richness and heterozygosity can be explained by
168 gene flow and phylogenetic structure, which is consistent with the expectation that the effect of gene
169 flow scales positively with the number of hybridising species and the divergence among these. In the,
170 an order of magnitude younger, cichlid radiation in Lake Malawi, heterozygosity levels vary much
171 less among lineages and do not scale with species richness, which – according to our findings – can
172 be explained by the much lower levels of genetic differentiation between the hybridising species³³.

173 **Conclusion**

174 Based on a comprehensive dataset on cichlid fishes from African Lake Tanganyika we tested
175 predictions related to the phenomenon of adaptive radiation. We establish that the Tanganyikan
176 cichlid radiation unfolded within the temporal and spatial confines of the lake, giving rise to an
177 endemic fauna consisting of ~240 species in 52 genera and 13 tribes in less than 10 Myr. Although
178 the ancestors of the tribes initially found comparable ecological opportunity, present-day species
179 numbers differ by two orders of magnitude among these phylogenetic sub-lineages. Our analyses of
180 morphological, ecological, and genomic information revealed that, taken as a whole, species-rich
181 tribes occupy larger fractions of the morphospace and ecospace and contain species that are, at the
182 per-genome level, genetically more diverse, which appears to be linked to gene flow. We demonstrate
183 a phenotype-environment association in three trait complexes (body shape, upper oral jaw
184 morphology, and lower pharyngeal jaw shape) and pinpoint their most relevant adaptive components.
185 Importantly, we show that eco-morphological diversification was not gradual over the course of the
186 radiation. Instead, we identified trait-specific pulses of accelerated phenotypic evolution, whereby
187 only diversification in body shape shows an ‘early burst’^{1,5}. The sequence of the trait-specific pulses
188 essentially follows the pattern postulated in the stages model of adaptive radiation^{1,6,7}, with the
189 extension that the most recent ‘stage’ of the cichlid adaptive radiation in Lake Tanganyika, which is
190 characterised by a large number of speciation events, is defined by increased diversification in both
191 an ecological (lower pharyngeal jaw) and a signalling (pigmentation) trait. To what extent the
192 observed diversity and disparity patterns were shaped by past environmental fluctuations and
193 extinction dynamics cannot be answered conclusively through the investigation of the extant fauna
194 alone.

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260 **Figure Legends**

261 **Fig. 1 | Time-calibrated species tree of the cichlid fish fauna of African Lake Tanganyika.** The species
262 tree was time calibrated with a relaxed-clock model and is based on a maximum-likelihood topology inferred
263 from genome-wide nuclear SNPs. Species names are abbreviated using a six-letter code, whereby the first
264 three letters represent the genus and the last three letters the species name (Extended Data Table 1; see
265 Extended Data Fig. 2a for the phylogeny with full species names). Branches are coloured according to tribes
266 and for all endemic species an illustration is shown. Representatives of riverine cichlids (grey font) are nested
267 within the radiation. The inset shows the time-calibrated phylogeny of more ancestral cichlid lineages
268 (estimated under the multi-species coalescent model, Extended Data Fig. 1), highlighting the phylogenetic
269 positions of the Tanganyikan representatives of the tribes Coptodonini (*Copren*; *Coptodon rendalli*),
270 Oreochromini (*Oretan*; *Oreochromis tanganyicae*), Tylochromini (*Tylpol*; *Tylochromis polylepis*) which
271 colonised the lake secondarily. The schematic map of the African continent shows the position of the three
272 Great Lakes Victoria, Malawi, and Tanganyika, with a magnified section of the latter. The presumed age of
273 Lake Tanganyika (9–12 Ma)²³ is indicated in blue along the time axes. Species trees based on alternative
274 topologies are presented in Extended Data Fig. 2b,c, and uncalibrated nuclear and mitochondrial phylogenies
275 on the specimen level are shown in Extended Data Fig. 3.

276 **Fig. 2 | Morphospace and ecospace occupation of the cichlid fish fauna of Lake Tanganyika.** Principal
277 component analyses of body shape (**a**, $n = 242$ taxa; 2,197 specimens), upper oral jaw morphology (**b**, $n = 242$;
278 2,197 specimens) and lower pharyngeal jaw shape (**c**, $n = 239$) along with the associated shape changes. **d**,
279 Ecospace spanned by the stable C and N isotope compositions ($\delta^{13}\text{C}$ and $\delta^{15}\text{N}$ values; $n = 236$; 1,168
280 specimens). The colour scale indicates the number of species in 20 by 20 bins across the morpho- and ecospace,
281 respectively (see Extended Data Fig. 5 for PCA and stable isotopes biplots with a focus on morpho- and
282 ecospace occupation per tribe).

283 **Fig. 3 | Temporal dynamics of diversification in body shape (first row), upper oral jaw morphology**
284 **(second row), lower pharyngeal jaw shape (third row), and pigmentation patterns (fourth row) in the**
285 **adaptive radiation of cichlid fishes in Lake Tanganyika.** **a**, Species tree (Fig. 1) with branches coloured
286 according to the mean relative rates of trait evolution for each trait. PP = posterior probability for rate shift. **b**,
287 Morphospace densities (number of lineages) through time for each trait. Blue lines indicate the point in time
288 when 50% of the extant morphospace had become occupied. **c**, Comparison of slopes (blue) of morphospace
289 expansion over time between the observed data and the Brownian motion null model of trait evolution. A
290 difference in slopes above zero represents morphospace expansion and values below zero indicate
291 morphospace packing relative to the null model. The shaded areas show 95% quantiles of the 500 Brownian
292 motion simulations. Lineage accumulation through time derived from the species tree is shown in dark grey.
293 **d**, Mean relative rates of trait evolution over time with standard deviation (blue). Lineage accumulation
294 through time is shown in dark grey.

295 **Fig. 4 | Association between genomic features and species richness across the cichlid tribes in Lake**
296 **Tanganyika.** Each genomic summary statistic was tested for a correlation with species richness per tribe (log-
297 transformed). To account for phylogenetic structure in the data, we calculated phylogenetic independent
298 contrasts for each variable. Data points are coloured according to tribes; large points are tribe means, shown
299 with 95% confidence intervals, small points represent species means and are only shown for group sizes < 40.

300 **a**, Percentage of the genome identified as transposable elements (TEs) (Pearson's $r = -0.31$, $df = 10$, $P = 0.33$;
301 tribe means are based on one genome per species; Extended Data Fig. 7a). **b**, Number of duplicated genes
302 (Pearson's $r = -0.27$, $df = 10$, $P = 0.40$; tribe means are based on species means). **c**, Genome-wide dN/dS ratios
303 as a measure of selection on coding sequences (Pearson's $r = 0.26$, $df = 10$, $P = 0.42$; tribe means are based on
304 species means across a set of 15,294 genes per genome; Extended Data Fig. 7b). **d**, Percentage of heterozygous
305 sites per genome (Pearson's $r = 0.70$, $df = 10$, $P = 0.012$; tribe means are based on species means). **e**, f_4 -ratio
306 statistics as a measure of gene flow among species within each tribe (Pearson's $r = -0.35$, $df = 9$, $P = 0.29$;
307 tribe means are based on all species triplets within each tribe; see Extended Data Fig. 8b for a summary of the
308 f_4 -ratio statistics for all species comparisons). **f**, Mean percentage of heterozygous sites in simulations with
309 within-tribe migration rates sampled from the observed f_4 -ratio statistics (Pearson's $r = 0.85$, $df = 10$, $P <$
310 0.001 ; tribe means are based on species means across 20 simulations; Extended Data Fig. 7c).

311 **Methods**

312 **Sampling**

313 Sampling was conducted between 2014 and 2017 at 130 locations at Lake Tanganyika. To maximise taxon
314 coverage, we included additional specimens from previous expeditions (4.9% of the samples) as well as from
315 other collections (0.8%). The final dataset (301 taxa; $n = 2,723$ specimens) contained an almost complete taxon
316 sampling of the cichlid fauna of Lake Tanganyika, as well as 18 representative cichlid species from nearby
317 waterbodies, and 32 outgroup species. All analyses described below are based on the same set of typically 10
318 specimens per species, or subsets thereof (see Extended Data Table 1 and Supplementary Methods for details).

319 **Whole genome sequencing**

320 Genomic DNA of typically one male and one female specimen per species ($n = 547$) was extracted from fin-
321 clips preserved in ethanol using the E.Z.N.A. Tissue DNA Kit (Omega Bio-Tek) and sheared on a Covaris
322 E220 (60 μ l with 10% duty factor, 175 W, 200 cycles for 65 sec). Individual libraries were prepared using
323 Illumina's TruSeq DNA PCR-Free Sample Preparation kit (Low Sample Protocol) for 350 bp insert size,
324 pooled (six per lane), and sequenced at 126 bp paired-end on an Illumina HiSeq 2500 (Extended Data Table 1
325 contains information on read depths).

326 **Assessing genomic variation**

327 After adapter removal with Trimmomatic³⁵ (v.0.36), reads of 528 genomes (all species belonging to the cichlid
328 radiation in Lake Tanganyika plus additional species nested within this radiation and some selected outgroup
329 species; Extended Data Table 1) were mapped to the Nile tilapia reference genome (RefSeq accession
330 GCF_001858045.1³⁶) using BWA-MEM³⁷ (v.0.7.12). Variant calling was performed with GATK's
331 HaplotypeCaller and GenotypeGVCF tools³⁸ (v.3.7), applying a minimum base quality score of 30. Variant
332 calls were filtered with BCFtools³⁹ (v.1.6; FS < 20, QD > 2, MQ > 20, DP > 4000, DP < 8000,
333 ReadPosRankSum > -0.5, MQRankSum > -0.5). We applied a filter to sites in proximity to indels with a minor
334 allele count greater than 2, depending on the size of the indel. With SNPable
335 (<http://lh3lh3.users.sourceforge.net/snpable.shtml>), we determined all sites within regions of the tilapia
336 reference genome in which read mapping could be ambiguous and masked these sites. Using VCFtools⁴⁰
337 (v.0.1.14) we further masked, per individual, genotypes with a read depth below 4 or a genotype quality below
338 20. Sites that were no longer polymorphic after the filtering steps were excluded, resulting in a dataset of
339 57,751,375 SNPs. Called variants were phased with the software beagle⁴¹ (v.4.1). The phasing of
340 *Neolamprologus cancellatus*, which appeared to be F1 hybrids, was further improved with a custom script.
341 Further details are provided in the Supplementary Methods.

342 **De novo genome assemblies**

343 *De novo* genome assemblies were generated from the raw-read data for each individual following an approach
344 described previously^{42,43}, using CeleraAssembler⁴⁴ (v.8.3) and FLASH⁴⁵ (v.1.2.11). Eight genomes repeatedly
345 failed to assemble and were therefore excluded from further analyses (specimen vouchers: A188, IRF6, IZC5,
346 JWE7, JWG1, JWG2, LJD3, and LJE8). Assembly quality was assessed with QUASt⁴⁶ (v.4.5) and
347 completeness was determined with BUSCO⁴⁷ (v.3). Assembly statistics summarised with MultiQC⁴⁸ (v.1.7)
348 are available on Dryad.

349 **Determining the age of the radiation**

350 To determine the age of the cichlid radiation in Lake Tanganyika, we applied phylogenomic molecular-clock
351 analyses for representatives of all cichlid subfamilies and the most divergent tribes, together with non-cichlid
352 outgroups (44 species; Extended Data Fig. 1). Following Matschiner et al.¹⁸ we identified and filtered ortholog
353 sequences from genome assemblies and compiled ‘strict’ and ‘permissive’ datasets that contained alignments
354 for 510 and 1,161 genes and had total alignment lengths of 542,922 and 1,353,747 bp, respectively. We first
355 analysed the topology of the species with the multi-species coalescent model implemented in ASTRAL⁴⁹
356 (v.5.6.3), based on gene trees that we estimated for both datasets with BEAST2⁵⁰ (v.2.5.0).

357 As undetected past introgression can influence divergence-time estimates in molecular clock analyses,
358 we further tested for signals of introgression in the form of asymmetric species relationship in gene trees and
359 excluded five species (*Fundulus heteroclitus*, ‘*Tilapia*’ *brevimanus*, *Pelmatolapia mariae*, *Tilapia sparrmanii*,
360 and *Steatocranus* sp. “ultraslender”) potentially affected by introgression from all subsequent molecular-clock
361 analyses. We then estimated divergence times among the most divergent cichlid tribes and the age of the
362 radiation with the multi-species coalescent model in StarBEAST2⁵¹ (v.0.15.5), using the ‘strict’ set of gene
363 alignments (Extended Data Fig. 1). Further details are provided in the Supplementary Methods.

364 **Phylogenetic inference**

365 To infer a complete phylogeny of the cichlid radiation in Lake Tanganyika (the Tanganyikan representatives
366 of the more ancestral tribes Coptodonini, Oreochromini, and Tylochromini were excluded) from genome-wide
367 SNPs we applied additional filters, retaining only SNPs with < 40% missing data and between-SNP distances
368 of at least 100 bp. The remaining 3,630,997 SNPs were used to infer a maximum-likelihood phylogeny with
369 RAxML⁵² (v.8.2.4; Fig. 1, Extended Data Fig. 2a, 3a). The species-tree topology was further estimated under
370 the multi-species coalescent model from a set of local phylogenies with ASTRAL (Extended Data Fig. 2b);
371 these local phylogenies were inferred with IQ-TREE⁵³ (v.1.7-beta7) from alignments for 1,272 genomic
372 regions determined to be particularly suitable for phylogenetic analysis (see Supplementary Methods). We also
373 applied the multi-species coalescent model implemented in SNAPP⁵⁴ (v.1.4.2) to the dataset of genome-wide
374 SNPs (Extended Data Fig. 2c). Species-level phylogenies resulting from these different approaches were used
375 as topological constraints in subsequent relaxed-clock analyses of divergence times (see below). In addition,
376 we estimated the mitochondrial phylogeny based on maximum-likelihood with RAxML. Further details are
377 provided in the Supplementary Methods.

378 **Divergence time estimates within the radiation**

379 For relaxed-clock analyses, the 1,272 alignments were further filtered by applying stricter thresholds on the
380 proportion of missing data and the strength of recombination signals. Ten remaining alignments with a length
381 greater than 2,500 bp and less than 130 hemiplasies (total length: 30,738 bp; completeness: 95.8%), were then
382 used jointly to estimate divergence times with the uncorrelated-lognormal relaxed-clock model implemented
383 in BEAST2. To account for phylogenetic uncertainty in downstream phylogenetic comparative analyses, we
384 performed three separate sets of relaxed clock analyses, in which the topology was either fixed to the species-
385 level phylogeny inferred with RAxML (Fig. 1, Extended Fig. 2a), the species tree inferred with ASTRAL

386 (Extended Fig. 2b), or the Bayesian species tree inferred with SNAPP (Extended Fig. 2c). Further details are
387 provided in the Supplementary Methods.

388 **Morphometrics**

389 To quantify body shape and upper oral jaw morphology we applied a landmark-based geometric morphometric
390 approach to digital X-ray images (for the full set of 10 specimens per species whenever possible; $n = 2,197$).
391 We selected 21 landmarks, of which 17 were distributed across the skeleton and four defined the premaxilla
392 (Extended Data Fig. 4a). Landmark coordinates were digitised using FIJI⁵⁵ (v2.0.0-rc-68/1.521i). To extract
393 overall body shape information, we excluded landmark 16, which marks the lateral end of the premaxilla,
394 hence minimizing the impact of the orientation of the upper oral jaw. We then applied a Procrustes
395 superimposition to remove the effect of size, orientation, and translational position of the coordinates.

396 For upper oral jaw morphology, we used a subset of four landmarks. A crucial feature of the oral jaw
397 morphology is the orientation of the mouth relative to the body axes. However, this component of the upper
398 oral jaw morphology would be lost in a classical geometric morphometric analysis, in which only pure shape
399 information is retained. To overcome this, we extracted the premaxilla-specific landmarks (1, 2, 16, and 21)
400 *after* Procrustes superimposition of the entire set of landmarks and subsequently re-centred the landmarks to
401 align the specimens without rotation. Thus, the resulting landmark coordinates do not represent the pure shape
402 of the premaxilla but additionally contain information on its orientation and size in relation to body axes and
403 body size, respectively.

404 To quantify lower pharyngeal jaw bone shape in 3D, a landmark-based geometric morphometric
405 approach was applied on μ CT-scans of the head region of five specimens per species ($n = 1,168$). To capture
406 all potential functionally important structures of the lower pharyngeal jaw bone, we selected a set of 27
407 landmarks (10 true landmarks and 17 sliding semi-landmarks) well distributed across the left side of the bone
408 (Extended Data, Fig. 4b). Landmark coordinates were acquired using TINA⁵⁶ (v.6.0). To retain the lateral
409 symmetric properties of the shape data during superimposition, we reconstructed the right side of the lower
410 pharyngeal jaw bone by mirroring the landmark coordinates across the plane of bilateral symmetry fitted
411 through all landmarks theoretically lying on this plane. We then superimposed the resulting 42 landmarks
412 while sliding the semi-landmarks along the curves by minimizing Procrustes distances and retained the
413 symmetric component only.

414 To identify the major axes of shape variation across the multivariate datasets we performed a PCA for
415 each trait. We also calculated morphospace size per tribe as the square root of the convex hull area spanned by
416 species means of the PC1- and PC2-scores. We then tested for a correlation between morphospace size and
417 estimated species richness of a tribe¹⁵ (log-transformed to obtain normal distribution). To account for
418 phylogenetic non-independence, we calculated phylogenetic independent contrasts with the R package *ape*⁵⁷
419 (v.5.2) using the species tree (Fig. 1) pruned to the tribe level. We then calculated Pearson's correlation
420 coefficients for independent contrasts using the function *cor.table* of the R package *picante*⁵⁸ (v.1.8).

421 All landmark coordinates for geometric morphometric analyses were processed and analysed in R⁵⁹
422 (v.3.5.2) using the packages *geomorph*⁶⁰ (v.3.0.7) and *Morpho*⁶¹ (v.2.6). Further details are provided in the
423 Supplementary Methods.

424 **Stable isotope analysis**

425 To approximate ecology for each species, we measured the stable carbon (C) and nitrogen (N) isotope
426 composition of all available specimens from Lake Tanganyika ($n = 2,259$). We analysed a small (0.5 – 1 mg)
427 dried muscle sample of each specimen with a Flash 2000 elemental analyser coupled to a Delta Plus XP
428 continuous-flow isotope ratio mass spectrometer (IRMS) via a ConFlo IV interface (Thermo Fisher Scientific,

429 Bremen, Germany). Carbon and nitrogen isotope data were normalised to the VPDB (Vienna Pee Dee
430 Belemnite) and Air-N₂ scales, respectively, using laboratory standards which were calibrated against
431 international standards. Values are reported in standard per-mil notation (‰), and long-term analytical
432 precision was 0.2‰ for δ¹³C values and 0.1‰ for δ¹⁵N values. To test for a correlation of ecospace size with
433 species richness of the tribes, we applied the same approach as described above to the δ¹³C and δ¹⁵N values.

434 To confirm interpretability of the δ¹³C and δ¹⁵N values, we additionally collected and analysed
435 baseline samples covering several trophic levels from the northern and the southern basin of Lake Tanganyika
436 (Supplementary Methods, Supplementary Discussion).

437 **Phenotype-environment association**

438 For each trait (body shape, upper oral jaw, lower pharyngeal jaw) we performed a two-block PLS analysis
439 based on species means of the Procrustes aligned landmark coordinates and the stable C and N isotope
440 compositions using the function *two.b.pls* in Geomorph. To account for phylogenetic dependence of the data
441 we applied a pGLS as implemented in the R package *caper*⁶² (v.1.0.1) across the two sets of PLS scores (each
442 morphological axis and the stable isotope projection) using the time-calibrated species tree based on the
443 maximum-likelihood topology. The strength of phylogenetic signal in the data was accounted for by optimising
444 the branch length transformation parameter lambda using a maximum-likelihood approach.

445 **Scoring pigmentation patterns**

446 To quantify a putative signalling trait in cichlids, we scored the pigmentation patterns in typically five male
447 specimens per species ($n = 1,028$), on the basis of standardised images taken in the field after capture of the
448 specimens (see Supplementary Methods). Following the strategy described in Seehausen et al.⁶³, the
449 presence/absence of 20 pigmentation features was recorded, whereby we extended number of scored features
450 to include additional body and fin pigmentation patterns (Extended Data Fig. 4c). We then applied a logistic
451 PCA implemented in the R package *logisticPCA*⁶⁴ (v.0.2) and used the PC1-scores as univariate proxy for
452 differentiation along the signalling axes for further analyses.

453 **Trait evolution modelling and disparity estimates**

454 To investigate the temporal dynamics of morphological diversification over the course of the radiation we
455 essentially followed the strategy of Cooney et al.²⁸ (which is based on measurements on extant taxa and
456 assumes constant niche-space and no or constant extinction over the course of the radiation), using the PLS-
457 scores of body shape, upper oral jaw morphology, and lower pharyngeal jaw shape and the PC1-scores of
458 pigmentation patterns as well as the time-calibrated maximum-likelihood species tree topology. For each trait
459 we assessed the phylogenetic signal in the data by calculating Pagel's Lambda and Blomberg's K with the R
460 package *phytools*⁶⁵ (v.0.6-60). We then tested the fit of four models of trait evolution for each of the four traits.
461 We applied a white noise model, a Brownian motion model, a single-optimum Ornstein-Uhlenbeck model,
462 and an 'early burst' model of trait evolution using the function *fitContinuous* of the R package *geiger*⁶⁶
463 (v.2.0.6.1). Additionally, we fitted a variable rates model (a Brownian motion model which allows for rate
464 shift on branches and nodes) using the software *BayesTrait* (<http://www.evolution.rdg.ac.uk/>, v.3) with
465 uniform prior distributions adjusted to our dataset (alpha: -1 – 1, sigma: 0 – 0.001 for morphometric traits;
466 alpha: 0 – 10, sigma: 0 – 10 for pigmentation pattern) and applying single-chain Markov-chain Monte Carlo
467 runs with one billion iterations. We sampled parameters every 100,000th iteration, after a pre-set burnin of
468 10,000,000 iterations. We then tested for each trait for convergence of the chain using a Cramer-von-Mises
469 statistic implemented in the R package *coda*⁶⁷ (v.0.19-3). The models were compared by calculating their log-
470 likelihood and Akaike information criterion (AIC) difference (Extended Data Table 2). Based on differences
471 in AIC, the variable rates model was best supported for all traits but body shape, which showed a strong signal
472 of an early burst of trait evolution (Extended Table 2, note that the variable rates model has the highest log-

473 likelihood for body shape as well). We nevertheless focused on the variable rates model for further analyses
474 of all traits to be able to compare temporal patterns of trait evolution among the traits.

475 To estimate morphospace expansion through time we used a maximum-likelihood ancestral state
476 reconstruction implemented in phytools. To account for differences in the rate of trait evolution along the
477 phylogeny, we reconstructed ancestral states using the mean rate-transformed tree derived from the variable
478 rates model. We then projected the ancestral states onto the original species tree and calculated the
479 morphospace extent (i.e. the range of trait values) in time intervals of 0.15 million years (note that this is an
480 arbitrary value; however, differently sized time intervals had no effect on the interpretation of the results). For
481 each time point we extracted the branches existing at that time and predicted the trait value linearly between
482 nodes. We then compared the resulting morphospace expansion over time relative to a null model of trait
483 evolution. We therefore simulated 500 datasets (PLS and PC1 scores) under Brownian motion given the
484 original species tree with parameters derived from the Brownian motion model fit to the original data. For each
485 simulated dataset we produced morphospace-expansion curves using the same approach as described above.
486 We then compared the slopes of our observed data with each of the null models by calculating the difference
487 of slopes through time (Fig. 3) using linear models fitted for each time interval with the two subsequent time
488 intervals. (Note that for body shape we also estimated morphospace expansion through time using the early
489 burst model for ancestral state reconstruction, which resulted in a very similar pattern of trait diversification.)

490 Unlike other metrics of disparity (e.g. variance or mean pairwise distances) morphospace extent is not
491 sensitive to the density distribution of measurements within the morphospace and captures its full range⁶⁸.
492 Hence, comparing the extent of morphospace between observed data and the null model directly unveils the
493 contribution of morphospace expansion relative to the null model; and because the increase in lineages over
494 time is identical in the observed and the simulated data, this comparison also provides an estimate for
495 morphospace packing.

496 To summarise evolutionary rates we calculated the mean rate of trait evolution inferred by the variable
497 rates model in the same 0.15 million years intervals along the phylogeny.

498 To account for phylogenetic uncertainty in the tree topology we repeated the analyses of trait evolution
499 using the time-calibrated trees based on tree topologies estimated with ASTRAL and SNAPP (Extended Data
500 Fig. 2b,c; Supplementary Methods; Supplementary Discussion). Furthermore, to also account for uncertainty
501 in branch lengths, we repeated the analysis on 100 trees from the Bayesian posterior distribution for each of
502 the three trees.

503 Further details can be found in the Supplementary Methods.

504 **Characterisation of repeat content**

505 For the repeat content analysis, we randomly selected one *de novo* genome assembly per species of the
506 radiation ($n = 245$). We performed a *de novo* identification of repeat families using RepeatModeler⁶⁹ (v.1.0.11).
507 We then combined the RepeatModeler output library with the available cichlid-specific libraries⁷⁰ (Dfam and
508 RepBase; v.27.01.2017; 258 ancestral and ubiquitous sequences, 161 cichlid-specific repeats, and 6 lineage-
509 specific sequences; 65,118, 273,530, and 6,667 bp in total, respectively) and used the software RepeatMasker⁷⁰
510 (v.4.0.7) (-xsmall -s -e ncbi -lib combined_libraries.fa) to identify and soft-mask interspersed repeats and low
511 complexity DNA sequences in each assembly. The reported summary statistics were obtained using
512 RepeatMasker's 'buildSummary.pl' script (Fig. 4a, Extended Data Fig. 7a, results per genome are provided
513 on Dryad).

514 **Gene duplication estimates**

515 Per genome, gene duplication events were identified with the structural variant identification pipeline smooove
516 (population calling method; <https://github.com/brentp/smoove>, docker image cloned 20/12/2018), which
517 builds upon lumpy⁷¹, svtyper⁷², and svtools (<https://github.com/hall-lab/svtools>). Variants were called per
518 sample ($n = 488$ genomes, 246 taxa of the Tanganyika radiation) from the initial mapping files against the
519 tilapia reference genome with the function *call*. The union of sites across all samples was obtained with the
520 function *merge*, then all samples were genotyped at those sites with the function *genotype*, and depth
521 information was added with *--duphold*. Genotypes were combined with the function *paste* and annotated with
522 *annotate* and the reference genome annotation file. The obtained VCF file was filtered with BCFtools to keep
523 only duplications longer than 1 kb and of high quality (MSHQ > 3 or MSHQ == -1, FMT/DHFFC[0] > 1.3,
524 QUAL > 100). The resulting file was loaded into R (v.3.6.0) with vcfR⁷³ (v.1.8.0) and filtered to keep only
525 duplications with less than 20% missing genotypes. Next, we removed duplication events with a length outside
526 1.5 times the interquartile range above the upper quartile of all duplication length, resulting in a final dataset
527 of 476 duplications (Fig. 4b).

528 **Analyses of selection on coding sequence**

529 To predict genes within the *de novo* genome assemblies, we used AUGUSTUS⁷⁴ (v.3.2.3) with default
530 parameters and ‘zebrafish’ as *--species* parameter ($n = 485$ genomes, 245 taxa). For each prediction we inferred
531 orthology to Nile tilapia genes (GCF_001858045.1_ASM185804v2) with GMAP (GMAP-GSNAP⁷⁵; v.2017-
532 08-15) applying a minimum trimmed coverage of 0.5 and a minimum identity of 0.8. We excluded specimens
533 with less than 18,000 tilapia orthologous genes detected ($n = 471$ genomes, 243 taxa). Next, we kept only those
534 tilapia protein coding sequences that had at least one of their exons present in at least 80% of the assemblies
535 (260,335 exons were retained, representing 34,793 protein coding sequences). Based on the tilapia reference
536 genome annotation file, we reconstructed for each assembly the orthologous coding sequences. Missing exon
537 sequences were set to ‘N’s. We then kept a single protein coding sequence per gene (the one being present in
538 the maximum number of species with the highest percentage of sequence length), resulting in 15,294 protein
539 coding sequences. Per gene, a multiple sequence alignment was then produced using MACSE⁷⁶ (v.2.01). We
540 calculated for each specimen and each gene the number of synonymous (S) and non-synonymous (N)
541 substitutions by pairwise comparison to the ortholog tilapia sequence using *codeml* with *runmode* -2 within
542 PAML⁷⁷ (v.4.9e). To obtain an estimate of the genome-wide sequence evolution rate that is independent of
543 filtering thresholds, we calculated the genome-wide dN/dS ratio for each specimen based on the sum of dS
544 and dN across all genes (Fig. 4c, Extended Data Fig. 7b).

545 **Signals of past introgression**

546 We used the f_4 -ratio statistic³⁴ to assess genomic evidence for interspecific gene exchange. We calculated the
547 f_4 -ratio for all combinations of trios of species on the filtered VCF files using the software Dsuite⁷⁸ (v.0.2 r20),
548 with *T. sparrmanii* as outgroup species (we excluded *N. cancellatus* as all specimens of this species appeared
549 to be F1 hybrids; Supplementary Methods). The f_4 -ratio statistic estimates the ‘admixture proportion’, i.e. the
550 proportion of the genome affected by gene flow. The results presented in this manuscript (Fig. 4e, Extended
551 Data Fig. 8a) are based on the ‘tree’ output of the Dsuite function *Dtrios*, with each trio arranged according to
552 the species tree based on the maximum-likelihood topology. The per-tribe analyses (Fig. 4e) were based only
553 on comparisons where all species within a trio belong to the same tribe ($n = 243$ taxa).

554 In addition to the f_4 -ratio we also identified signals of past introgression among species using a
555 phylogenetic approach by testing for asymmetry in the relationships of species trios in 1,272 local maximum-
556 likelihood trees generated using IQ-TREE (Supplementary Methods; Extended Data Fig. 8b).

557 **Heterozygosity**

558 We calculated the number of heterozygous sites per genome ($n = 488$ genomes, 246 taxa from the Tanganyika
559 radiation) from the VCF files using the BCFtools function *stats* and then quantified the percentage of
560 heterozygous sites among the number of callable sites per genome (see above) (Fig. 4d).

561 To explore if the observed levels of heterozygosity per tribe can be explained by the levels of gene
562 flow within tribes we performed coalescent simulations with msprime⁷⁹ (v.0.7.4). We simulated genome
563 evolution of all species of the radiation following the time-calibrated species tree (maximum-likelihood
564 topology), assuming a generation time of 3 years⁸⁰ and a constant effective population size of 20,000
565 individuals. Species divergences were implemented as mass migration events and introgression within tribes
566 as migration between species pairs with rates set according to their introgression (f_4 -ratio) signals inferred with
567 Dsuite. To convert the f_4 -ratio values into migration rates, we applied a scaling factor of 5×10^{-6} , which results
568 in a close correspondence in magnitude of the simulated introgression signals to those observed empirically
569 (Fig. 4, Extended Data Fig. 7c). In each of twenty separate simulations, we randomly sampled one pairwise f_4 -
570 ratio value for each pair of species (there are many f_4 -ratios per species pair – one for each possible third
571 species added to the test trio; the maximum values per pair are shown in Extended Data Figure 8a). The
572 simulated data consisted of one chromosome of 100 kb (mutation rate: 3.5×10^{-9} per bp per generation³³,
573 recombination rate: 2.2×10^{-8} per bp per generation; see Supplementary Methods). Levels of heterozygosity
574 were calculated for all simulated datasets as described for the empirical data.

575 To account for between-tribe gene flow we further performed simulations in which migration between
576 tribes was also sampled from the empirical f_4 -ratio distribution. For simplicity in setting up the simulation
577 model, we assume that gene flow between tribes is ongoing until present day, which is clearly an overestimate
578 (see Supplementary Discussion). Nevertheless, the results of these simulations support our hypothesized
579 scenario, confirming that much of the variation in heterozygosity as well as its correlation with species richness
580 can be explained by the observed levels of gene flow.

581 **Correlation of genome-wide statistics with species richness**

582 We tested for a correlation between tribe means (based on species means) of each genomic summary statistics
583 (TE counts, number of gene duplications, genome-wide dN/dS ratio, per-genome heterozygosity, and f_4 -ratio,
584 as well as the heterozygosity and f_4 -ratio statistics derived from simulated genome evolution) and species
585 richness of the tribes, applying the same approach as described above for tests of correlation between morpho-
586 and ecospace size and species richness.

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726 **Author Contributions**

727 F.R., A.I., and W.S. designed this study (with input from H.H.B., A.K., and S.J.). F.R., A.I., H.H.B, and W.S.
728 collected the specimens in the field. F.R. and A.Böh. extracted DNA and prepared the libraries for sequencing.
729 S.J. coordinated sequencing. M.Mat. performed the mapping, variant calling, phylogenetic analyses, and
730 coalescent simulations. M.Mal. contributed to the variant calling pipelines and performed the f_4 -ratio statistics.
731 A.Böh. assembled the genomes and quantified gene duplications, A.E. conducted the dN/dS analyses, and
732 V.R. analysed transposable elements. A.Boi. assessed stable isotope compositions, H.H.B. radiographed the
733 specimens, and W.S. scored pigmentation patterns. F.R. performed μ CT-scanning, geometric morphometric
734 analyses, and all analyses incorporating morphological and ecological data. F.R. and W.S. wrote the
735 manuscript with contributions and/or feedback from all authors. All authors read and approved the final version
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737 **Additional Information**

738 The authors declare no **Competing interests**.

739 **Supplementary Information** is available for this paper.

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741 (walter.salzburger@unibas.ch).

742 **Data availability**

743 All newly sequenced genomes for this study and their raw reads are available from NCBI under the BioProject
744 accession number PRJNA550295 (<https://www.ncbi.nlm.nih.gov/bioproject/>). The VCF file, tree files,
745 summary statistics of the assembled genomes, and phenotypic datasets generated and analysed during this
746 study are available as downloadable files on Dryad ([https://datadryad.org/stash/share/13fM-](https://datadryad.org/stash/share/13fM-BDssqlCELXWdSgpnkeawCuOFvo-tA9o-vEiZ_k)
747 [BDssqlCELXWdSgpnkeawCuOFvo-tA9o-vEiZ_k](https://datadryad.org/stash/share/13fM-BDssqlCELXWdSgpnkeawCuOFvo-tA9o-vEiZ_k)).

748 **Code availability**

749 Code used to analyse the data is available on GitHub (https://github.com/cichlidx/ronco_et_al), except for
750 analyses where single commands from publicly available software were used and where all settings are fully
751 reported in the Methods and/or Supplementary Methods sections.

752 **Extended Data**

753 **Extended Data Fig. 1 | Age of the adaptive radiation of cichlid fishes in African Lake Tanganyika.** Time-
754 calibrated species tree of species representing divergent tribes and subfamilies within cichlids as well as
755 closely-related non-cichlid outgroups, generated with the multi-species coalescent model in StarBEAST2.
756 Nodes marked with a black dot were constrained according to species-tree analyses with ASTRAL. Node bars
757 indicate 95% highest-posterior density age intervals. Outgroup divergence times are not drawn to scale. Insets
758 visualise the prior distribution applied for the age of African cichlids according to Matschiner et al.¹⁸, as well
759 as posterior age estimates for Oreochromini and the cichlid adaptive radiation in Lake Tanganyika.

760 **Extended Data Fig. 2 | Alternative time-calibrated species trees of the cichlid adaptive radiation in Lake**
761 **Tanganyika.** Time-calibrated species tree based on **a**, the maximum-likelihood topology estimated with
762 RAxML (Fig. 1), **b**, the topology estimated with ASTRAL, and **c**, the topology estimated with SNAPP. Time
763 calibration used a relaxed-clock model in BEAST2, applied to a selected set of alignments with a total length
764 of 30,738 bp.

765 **Extended Data Fig. 3 | Individual-level phylogenies for the cichlid adaptive radiation in Lake**
766 **Tanganyika.** **a**, Maximum-likelihood tree inferred from nuclear SNPs. Node labels indicate the proportion of
767 data subsets supporting a clade (equal to 1 for all nodes without labels). **b**, Mitochondrial phylogeny inferred
768 from mitochondrial genomes using maximum-likelihood. Node labels represent bootstrap values (100 for all
769 nodes without labels).

770 **Extended Data Fig. 4 | Phenotyping of the specimens.** **a**, Two-dimensional landmarks placed on X-ray
771 images of the specimens. To quantify overall body shape we excluded landmark 16 (to minimise the effect of
772 the orientation of the oral jaw). To analyse upper oral jaw morphology we used landmarks 1, 2, 16, and 21. **b**,
773 Three-dimensional landmarks used to analyse lower pharyngeal jaw shape on μ CT-scans of the heads. True
774 landmarks are indicated in red, sliding semi-landmarks are indicated in blue. **c**, Body regions scored for
775 presence/absence of pigmentation patterns.

776 **Extended Data Fig. 5 | Morphospace and ecospace of the cichlid adaptive radiation in Lake Tanganyika.**
777 Scatter plots for each focal tribe (indicated with colours, see Fig. 1 for colour key) against the total morpho-
778 and ecospace (grey). Species ranges are indicated with convex hulls. **a-c**, PC1 and PC2 of body shape, upper
779 oral jaw morphology, and lower pharyngeal jaw shape, respectively. For shape changes associated with the
780 respective PC-axis see Fig. 2 and for full species names see Extended Data Table 1. **d**, Stable C and N isotope
781 compositions ($\delta^{15}\text{N}$ and $\delta^{13}\text{C}$ values). The additional plot shows $\delta^{15}\text{N}$ and $\delta^{13}\text{C}$ values of a baseline dataset
782 which confirms the interpretability of the stable isotope N and C composition in Lake Tanganyika (see
783 Supplementary Methods and Discussion). The last plot for each trait shows the size of the morpho- and

784 ecospace per tribe in relation to species numbers (body shape: Pearson's $r = 0.91$, $df = 9$, $P < 0.001$; upper oral
785 jaw morphology: Pearson's $r = 0.88$, $df = 9$, $P < 0.001$; lower pharyngeal jaw shape: Pearson's $r = 0.83$, $df =$
786 9 , $P = 0.002$; stable isotopes: Pearson's $r = 0.88$, $df = 9$, $P < 0.001$). Morphospace size was calculated as the
787 square root of the convex hull area spanned by species means.

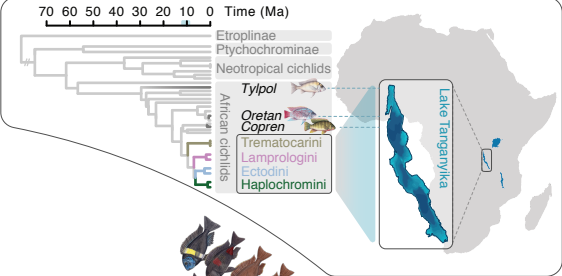
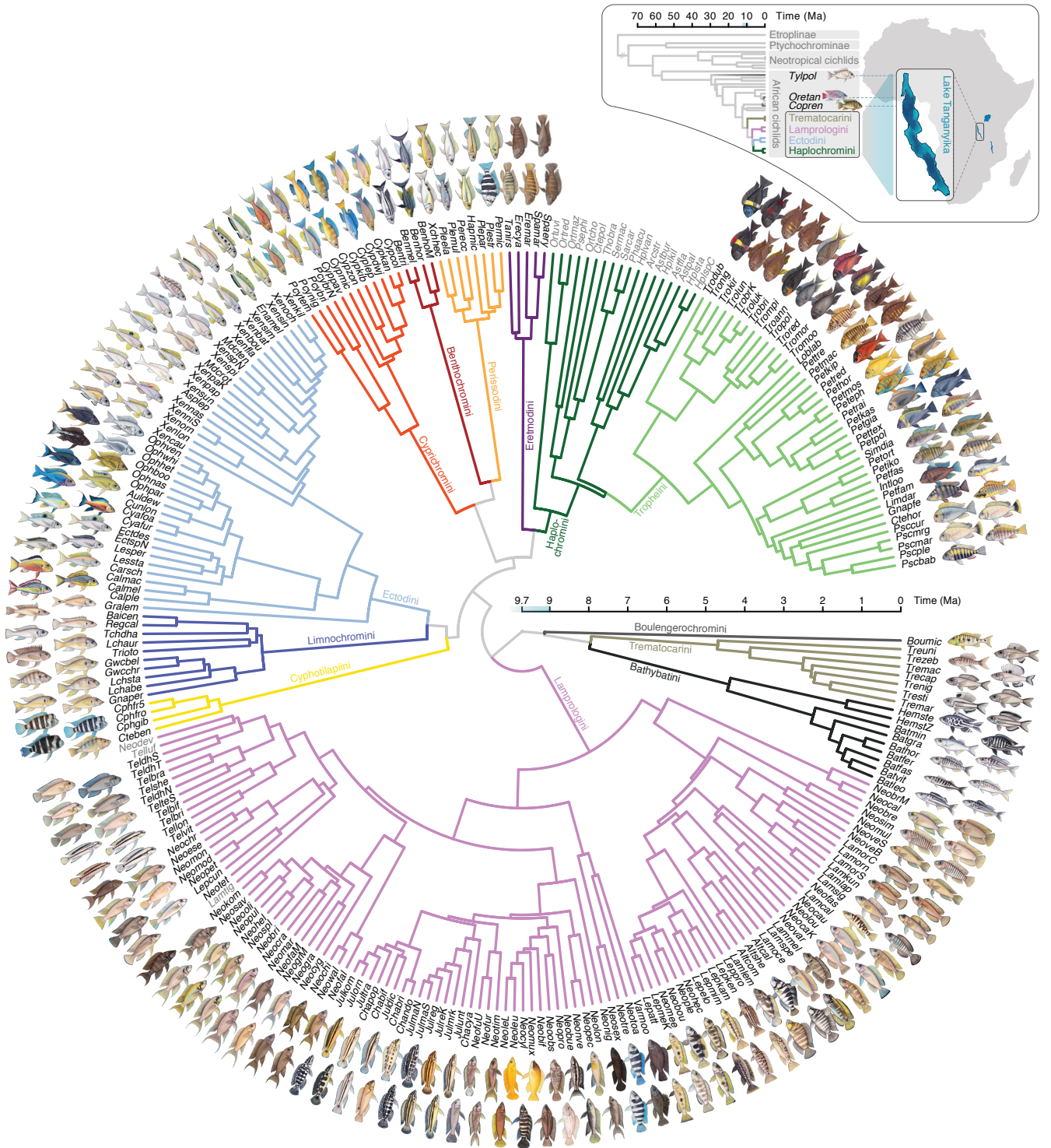
788 **Extended Data Fig. 6 | PLS fit for each multivariate trait against the stable C and N isotope compositions**
789 **($\delta^{15}\text{N}$ and $\delta^{13}\text{C}$ values).** Associated shape changes and loadings of the respective stable isotope projection are
790 indicated next to the axes. Data points represent species means and are coloured according to tribe. **a**, Body
791 shape; **b**, Upper oral jaw morphology; **c**, Lower pharyngeal jaw shape.

792 **Extended Data Fig. 7 | Genome-wide statistical analyses.** **a**, Proportion of the different classes of
793 transposable elements (TE) among all TE for each tribe (one genome per species). **b**, Species means of dN
794 (left) and dS (right) values over alignment length for each tribe. The boxes' centre lines show median, box
795 limits show first and third quartiles, and whiskers show the $1.5\times$ interquartile ranges. **c**, f_4 -ratio statistics among
796 species within each tribe in simulated data (tribe means are based on the mean across 20 simulations of each
797 species triplet). Data points are coloured according to tribes; large points are tribe means, shown with 95%
798 confidence intervals, small points represent species means and are only shown for group sizes < 40 species.
799 To test for a correlation with species richness per tribe (log-transformed), we first calculated phylogenetic
800 independent contrasts for each variable.

801 **Extended Data Fig. 8 | Signals of introgression among Lake Tanganyika cichlid species.** **a**, Maximum
802 values of the f_4 -ratio statistics between all pairs of species, derived from calculations across all combinations
803 of species trios with *Tilapia sparrmanii* fixed as the outgroup. The f_4 -ratio estimates the proportion of the
804 genome affected by gene flow, all presented values are statistically significant ($P < 5\times 10^{-5}$ after Benjamini-
805 Hochberg correction for multiple testing). **b**, D_{tree} -statistics (hue) with corresponding P -value (log-
806 transformed; saturation) based on a phylogenetic approach testing for asymmetry in the relationships of species
807 trios in 1,272 local maximum-likelihood trees (see Supplementary Methods). The two different approaches
808 uncovered little gene flow among the tribes (see Supplementary Discussion).

809 **Extended Data Table 1 | Sample size information per species.** For each analysis the total sample size is
810 given whereas the number in brackets indicates the number of specimens used uniquely for the respective
811 analysis. All genomes and raw sequences are available at NCBI under the BioProject accession number
812 PRJNA550295. A full list of individual specimen vouchers including details on sampling location is provided
813 as Supplementary Table 1. AMNH = American Museum of Natural History (New York, USA); MRAC =
814 Royal Museum for Central Africa (Tervuren, Belgium); HHB = Private collection of one of the authors, H.H.B.

815 **Extended Data Table 2 | Models of trait evolution.** **a**, Comparison of model fits for different models of trait
816 evolution and phylogenetic signal for each trait complex using three time-calibrated species trees with
817 alternative topologies. **b**, Overview of the model fits and phylogenetic signal inferred using 100 trees sampled
818 from the posterior distributions of the time calibrations for each of the three alternative tree topologies.



9.7 9 8 7 6 5 4 3 2 1 0 Time (Ma)

Boulengerochromini

Trematocarini

Bathybatini

Lamprologini

Boumic

Treuni

Trezeb

Trecap

Treng

Tresti

Tremar

Hemste

HemstZ

Batmin

Batgia

Batho

Batfas

Batle

Neobrt

Neobrs

Neobrn

NeoveS

LamoriC

LamoriS

Lamisp

Neosis

Neosac

Neosou

NeosouK

Neosou

Neosou

Neosou

